Assessment of Obesity and Fear of Fatness Among Inner-City Dublin Schoolchildren in a One-Year-Follow-Up Study

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Assessment of obesity and fear of fatness among inner-city Dublin schoolchildren in a one-year follow-up study

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Abstract

Background: Positive secular trends in adolescent obesity and an increased prevalence of fear of fatness, particularly among girls, have been documented world-wide. There is a lack of consensus about assessment criteria for childhood obesity and no standard exists for assessing Irish children. In 1990, the Irish National Nutrition Survey used body mass index (BMI) ≥ 26 kg m⁻² to describe the prevalence of overweight among Irish adolescents.

Objectives: (1) To examine the range in classification of Dublin schoolchildren as overweight according to four standard assessment methods; (2) to assess changes in weight status, prevalence of fear of fatness and accompanying slimming practices in a one-year follow-up; and (3) to compare the prevalence of overweight with that documented in 1990 among adolescents of similar age.

Design: A one-year follow-up study of 199 healthy schoolchildren (90 boys and 109 girls; mean age of 11 years at baseline) attending seven fee-paying (six single- and one mixed-sex) and eight non-fee-paying (four single- and four mixed-sex) primary schools in Dublin city centre.

Measurements: Weight, height, waist circumference and triceps skinfold were measured and used in five definitions of overweight, including published cut-off points of BMI-for-age (Centers for Disease Control and Prevention BMI-for-age charts for boys and girls; BMI reference curves for the UK 1990; International Obesity Task Force age- and sex-specific BMI cut-offs), actual relative weight and BMI ≥ 26 kg m⁻². Assessment of body image perceptions and satisfaction (using figure line drawings) was reported in a questionnaire specifically designed for this study.

Results: The prevalence of overweight within the total group differed between the four standard definitions of weight status, by 9% at baseline and 8% at follow-up. Accordingly, increasing trends over the year ranged from zero to 3%. Using the criterion BMI ≥ 26 kg m⁻², 6% of Dublin schoolchildren were overweight, compared with 1.9% of schoolchildren in 1990. Significantly more girls than boys were affected by fear of fatness and were trying to lose weight.

Conclusion: A standard method for assessment of weight status is urgently needed for the evaluation of obesity prevention initiatives among Irish schoolchildren. Such initiatives need to be sensitive to the pervasiveness of fear of fatness among adolescent girls.

The prevalence of childhood obesity is described as being of epidemic proportions around the world. The same serious long-term health problems caused by adult obesity are now found among overweight adolescents, fuelling the public health crisis. There has been no new information on body weight status among young Irish schoolchildren since the Irish National Nutrition Survey (INNS) in 1990, when a cut-off of body mass index (BMI) at 26 kg m⁻² identified 1.9% of schoolchildren (1.1% of boys and 2.6% of girls) aged 12–15 years as being at risk of overweight.

Assessments of the prevalence and trends in childhood overweight have been difficult to conduct owing to a lack of international consensus on the criteria recommended for classifying individuals as overweight or obese. Triceps skinfold closely relates to percentage body fat but is impracticable for epidemiological use. Weight-for-height indices are the most feasible and recently BMI has emerged as a means of defining overweight among children. Reference curves of BMI through childhood, that account for the substantial age-related changes, are...
necessary to assess BMI measurements. Current published BMI-for-age cut-offs include those of the Centers for Disease Control and Prevention (CDC), the BMI reference curves for the UK 1990 (UK90 charts) and the International Obesity Task Force (IOTF). Whereas the CDC in the USA and the UK90 charts in Great Britain are centile charts derived from national cross-sectional growth studies, the IOTF cut-offs are an internationally applicable standard derived from six large, nationally representative, cross-sectional growth studies. Obesity is defined as a BMI-for-age above the 95th centile according to the CDC and above the 98th centile according to the UK90 charts. These and the relevant IOTF cut-offs that correspond to 11- and 12-year-old boys and girls are presented in Table 1. The concern with having a variety of cut-offs for assessing overweight in the same population is that the prevalence of overweight will vary according to the reference applied. There are no BMI cut-offs representative of an Irish population. Irish growth charts are currently limited to overweight and actual overweight among children and adolescents.

BMI does not describe abdominal adiposity. Abdominal distribution of body fat and cardiovascular risk factors track from childhood to adulthood, and the identification of children with high abdominal adiposity is advantageous. Recent studies in children show that a greater distribution of body fat in the abdomen is correlated with less favourable patterns of serum lipoprotein concentrations and blood pressure. Recently, waist circumference has been shown to be a superior indicator of abdominal adiposity among children.

Concurrent with the increasing prevalence of obesity among children is a fear of fatness, particularly among young adolescent girls. Unhealthy eating and disordered eating behaviours among young adolescents have the potential to adversely affect nutrient intake, mental health status and long-term health outcomes. Compared with those who do not diet, adolescent girls who diet at a severe level are 18 times more likely, and those who diet at a moderate level five times more likely, to develop an eating disorder.

The purposes of this study were threefold. First, to examine the range in classification of Dublin schoolchildren as overweight according to four standard assessment methods, namely three published cut-off points of BMI-for-age (CDC, IOTF, UK90) and ARW. Second, to assess changes in weight status, prevalence of fear of fatness and accompanying slimming practices in a one-year follow-up. Finally, to indicate trends over the past decade, the prevalence of children with BMI ≥ 26 kg m⁻² was compared with that documented in the INNS among adolescents of similar age.

**Subjects and methods**

**Subjects**

Ethical approval for this research was obtained from the Joint Research Ethics Committee of Saint James’s Hospital and the Federated Dublin Voluntary Hospitals. A list of Dublin schools was obtained from the Department of Education and Science and those located in Dublin inner city were identified. Those that catered for special needs or the Irish language were excluded. Of 21 Dublin primary schools invited, 71% (n = 15) agreed to participate which included seven fee-paying (FP; six single- and one mixed-sex) and eight non-fee-paying (NFP; four single- and four mixed-sex) schools. Four hundred and seventy-six subjects, a response rate of 60%, received parental consent. A representative random sample (42% of the children recruited; n = 199, 90 boys, 109 girls; 88 FP, 111 NFP; mean age of 11 years at baseline), controlled equally for sex and school type attended, participated in the study.

**Socio-economic class**

Socio-economic class (SEC) was classified according to the highest classed earner (usually a parent or guardian) in the household.

**Anthropometry**

Two children at a time were measured in a private room at their school. Weight (to the nearest 0.1 kg) of each subject was measured barefoot and in light indoor clothing using

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**Table 1** The range in published cut-off points (either a given number or derived from reference centile charts), based on weight-for-height measurements, that are currently advised for defining risk of overweight and actual overweight among children and adolescents.

<table>
<thead>
<tr>
<th>Weight-for-height index</th>
<th>Risk of overweight</th>
<th>Overweight</th>
</tr>
</thead>
<tbody>
<tr>
<td>ARW (%)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>2–18 years</td>
<td>111–120</td>
<td>&gt; 120</td>
</tr>
<tr>
<td>Males: age 11 years</td>
<td>20.2</td>
<td>23.1</td>
</tr>
<tr>
<td>Females: age 11 years</td>
<td>20.8</td>
<td>24.1</td>
</tr>
<tr>
<td>Males: age 12 years</td>
<td>21.0</td>
<td>24.3</td>
</tr>
<tr>
<td>Females: age 12 years</td>
<td>21.7</td>
<td>25.2</td>
</tr>
<tr>
<td>UK90 (kg m⁻²)†</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Males: age 11 years</td>
<td>20.1</td>
<td>22.6</td>
</tr>
<tr>
<td>Females: age 11 years</td>
<td>21.2</td>
<td>23.8</td>
</tr>
<tr>
<td>Males: age 12 years</td>
<td>20.8</td>
<td>23.4</td>
</tr>
<tr>
<td>Females: age 12 years</td>
<td>21.9</td>
<td>24.8</td>
</tr>
<tr>
<td>IOTF (kg m⁻²)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Males: age 11 years</td>
<td>20.6</td>
<td>25.1</td>
</tr>
<tr>
<td>Females: age 11 years</td>
<td>20.7</td>
<td>25.4</td>
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<tr>
<td>Males: age 12 years</td>
<td>21.2</td>
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<tr>
<td>Females: age 12 years</td>
<td>21.7</td>
<td>26.7</td>
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</tbody>
</table>

ARW – actual relative weight; CDC – Centers for Disease Control and Prevention body mass index (BMI)-for-age charts for boys and girls; UK90 – BMI reference curves for the UK 1990; IOTF – International Obesity Task Force age- and sex-specific BMI cut-offs.

* Risk of overweight taken as ≥ 85th centile and < 95th centile; overweight taken as ≥ 95th centile.
† Risk of overweight taken as ≥ 91st centile and < 98th centile; overweight taken as ≥ 98th centile.
SECA weighing scales. Height (to the nearest 0.1 cm) was measured using a Leicester Height Measure (CMS Equipment). ARW was calculated as observed body weight expressed as a percentage of ideal body weight for height, age and sex, using Irish growth charts\textsuperscript{10}, and BMI was calculated from observed weight and height (kg m\textsuperscript{-2}). ARW cut-offs to define underweight, normal weight and overweight in this study were set at \(\leq 89\%\), 90–110\% and \(\equiv 111\%\), respectively.\textsuperscript{11} Table 1 gives the range of published cut-offs available to define the risk of overweight and actual overweight based on weight-for-height measurements in children at mean ages of 11 and 12 years. The proportions of children in the follow-up year (mean age 12 years) and children from the INNS\textsuperscript{3} (age range 12–15 years) who had BMI \(\leq 26\) kg m\textsuperscript{-2} were compared with the aim to elucidate the trend since 1990.

Waist circumference (WC; to the nearest mm) was measured at the minimum girth between the lower rib margin and the iliac crest over the naked site. Triceps were measured (to the nearest 0.1 mm) at the mid-point of the back of the upper right arm\textsuperscript{20} using a Harpenden Skinfold Calliper (British Indicators Ltd). The mean of three measurements, which did not differ by more than 1.0 mm, was used in analysis. In some cases an adequate measure of skinfold could not be recorded due to the tightness of the subcutaneous fat to muscle (6\% \(n = 5\) boys and 11\% \(n = 12\) girls at age 11 years, and 8\% \(n = 7\) boys and 15\% \(n = 16\) girls at age 12 years).

**Body image concerns**

Subjects completed a self-report questionnaire (approximately 20 minutes), specifically designed for this study, at their schools in an exam-type setting. The main investigator explained each question separately using an overhead projector and a standard script to ensure consistency and avoid possible misunderstanding of any question being asked. All questions included a ‘Don’t know’ option. Body image concerns were recorded using sex-specific line drawings\textsuperscript{21} illustrating seven body shapes ranging from severely underweight to severely overweight. The set of drawings applied in the current study was taken from an earlier study\textsuperscript{21} of body shape perceptions and preferences among 1118 preadolescent children (average age 8 years). In their development, the line drawings of children’s figures were reviewed by child and adult jurors, pilot tested, and examined for test–retest reliability and criterion validity\textsuperscript{21}. Before data collection for the present study, the self-report questionnaire was tested–retested among Dublin schoolchildren \((n = 34\textsuperscript{,} 17\text{ male, 17 female; mean age} 10.4\text{ years; from a mixed, NFP school) not participating in this study. A high proportion (82\%) gave consistent answers to the body image and simulating practices questions on retest, indicating reliability.\textsuperscript{21}} Subjects circled the line drawing that corresponded to the answer they associated with the question asked. Accordingly, disparity of own body shape was identified when the perceived body shape (‘Which one looks like you?’) line drawing circled by a subject differed from their measured ARW category. Circling an identical line drawing for preferred body shape (‘Which shape would you like to look like?’) as that chosen for perceived body shape (‘Which one looks like you?’) indicated satisfaction with body shape. Circling a thinner or fatter figure for preferred body shape than that chosen for perceived body shape indicated dissatisfaction with a preference towards being thinner or fatter, respectively. The subjects were also asked to report if they had ever tried to lose weight and which weight reduction methods they had used\textsuperscript{22}.

**Data management and statistical analysis**

All data were coded for entry into the Statistical Package for Social Sciences (SPSS) version 10.0 for Windows (SPSS Inc., Chicago, IL, USA). ‘Don’t know’ and ‘None of the above’ answers were omitted from analysis. Variables that were not normally distributed (using Kolmogorov–Smirnov tests), which included BMI, triceps and WC, were transformed (by taking the reciprocal or log) to achieve a normal distribution. O’Hares\textsuperscript{19} social classes one and two, three and four, five and six were combined to form socio-economic classes one, two and three, respectively. The seven categories of body shape illustrated by the line drawings were combined to define three distinct groups: underweight, normal weight and overweight.

Comparisons for the total group and for each sex were made between baseline and follow-up using the following statistics as appropriate: Wilcoxon signed rank, paired \(\text{t}-\) test and McNemar. Variables were analysed according to sex and SEC using the chi-square, Mann–Whitney and independent \(\text{t}-\) test, as appropriate. Pearson’s correlation was used to assess the strength of relationships between BMI and triceps skinfold. The proportions having BMI \(\leq 26\) kg m\textsuperscript{-2} and sample numbers from the current study and the INNS\textsuperscript{3} were used to calculate 95\% confidence intervals and difference in proportions. A significant difference was indicated by a \(P\)-value of \(\leq 0.05\).

**Results**

Anthropometric measurements were taken on average 0.98 (standard deviation (SD) 0.1) years (range 0.75–1.23 years) apart. The mean (SD) weight and height among boys were 41.1 (8.2) kg and 1.50 (0.07) m at age 11 years; and 46.7 (9.88) kg and 1.57 (0.08) m at age 12 years. The mean (SD) weight and height among girls were 42.4 (8.5) kg and 1.49 (0.07) m at age 11 years; and 48.2 (9.2) kg and 1.55 (0.07) m at age 12 years. The height gained by the boys was significantly greater than that gained by the girls (0.07 (0.02) m vs. 0.06 (0.02) m, \(P \leq 0.05\)). The mean (SD) for WC at mean age of 11 and 12 years was 64.5 (1.1) cm and 66.5 (1.1) cm, respectively, among boys; 64.3 (1.1) cm and 65.4 (1.1) cm among girls. The mean (SD) for triceps at
mean age of 11 and 12 years was, respectively, 10.9 (0.2) mm and 11.5 (0.2) mm among boys; 13.4 (0.1) mm and 13.9 (0.1) mm among girls. Triceps measurements among girls were significantly greater \((P \leq 0.001)\) than among boys at baseline and in the one-year follow-up. Anthropometric measurements did not differ according to SEC classification. At baseline and follow-up, BMI was found to correlate more strongly than ARW to triceps measurements \(\left( r = -0.731 \mathrm{vs.} \ r = 0.606 \right)\) at mean age of 11 years; \(r = -0.662 \mathrm{vs.} \ r = 0.581 \) at mean age of 12 years. (It should be noted that the distribution of BMI was skewed and was transformed to give a normal distribution by taking the reciprocal. Therefore, the negative correlation of BMI to triceps demonstrates that as BMI increases so does adiposity.)

The prevalence of overweight varied according to the different criteria currently available, as shown in Fig. 1a for mean age of 11 years and Fig. 1b for mean age of 12 years. The prevalence of overweight within the total group differed between criteria by 9% at baseline and 8% at follow-up. Positive trends in overweight prevalence in the one-year follow-up therefore ranged from zero to 3%. Using the criterion of the INNS \((\text{BMI} \geq 26 \text{ kg m}^{-2})\), it was found that 6% \((n = 12, 95\% \text{ CI} 0.5-4.8\%)\) of the total group, mean age 12 years, were overweight compared with 1.9% \((n = 4, 95\% \text{ CI} 0.3-7.9\%)\) of the total group \((n = 207)\) in 1990. This indicates a threefold increase \((95\% \text{ CI} 0.3-7.9\%), P \leq 0.05\) among 12-year-old children over the last 10 years. The prevalence of \(\text{BMI} \geq 26 \text{ kg m}^{-2}\) did not differ according to SEC classification.

According to ARW, overweight subjects had greater central adiposity than their normal-weight counterparts at mean age of 11 years \((71.2 (1.1) \text{ cm vs.} 63.2 (1.1) \text{ cm})\), \(P \leq 0.001\) and mean age of 12 years \((71.5 (1.1) \text{ cm vs.} 65.1 (1.1) \text{ cm})\), \(P \leq 0.001\). A greater proportion of boys than of girls were found to be underweight \((\text{at mean age of 11 years: } 36\% \ (n = 32) \mathrm{vs.} \ 25\% \ (n = 27), P \leq 0.05\); at mean age of 12 years: \(44\% \ (n = 40) \mathrm{vs.} \ 20\% \ (n = 22), P \leq 0.001\) at baseline and follow-up. Likewise, a smaller proportion of boys compared to girls were found to be overweight \((\text{at mean age of 11 years: } 17\% \ (n = 15) \mathrm{vs.} \ 32\% \ (n = 35), P \leq 0.05\); at mean age of 12 years: \(21\% \ (n = 19) \mathrm{vs.} \ 30\% \ (n = 33), P \leq 0.001\)\). The occupancy of ARW categories was independent of SEC classification.

According to ARW, approximately half of the total group \((47\% \ (n = 94) \mathrm{and} \ 51\% \ (n = 102)\) at mean age of 11 and 12 years, respectively) perceived their actual body shapes incorrectly (Fig. 2). Over a quarter of the children expressed a preference to be thinner \((29\% \ (n = 57) \mathrm{at} \ \text{mean age of 11 years}; 17\% \ (n = 15) \mathrm{vs.} \ 32\% \ (n = 35), P \leq 0.05\); at mean age of 12 years: \(21\% \ (n = 19) \mathrm{vs.} \ 30\% \ (n = 33), P \leq 0.001\)\). Preference towards thinness did not differ according to SEC classification. At baseline and follow-up respectively, overweight children \((58\% \ (n = 29) \mathrm{and} \ 54\% \ (n = 26)\) were more likely to express a desire for a thinner shape \((P \leq 0.001)\) than their normal-weight \((21\% \ (n = 19) \mathrm{and} \ 32\% \ (n = 26)\) and underweight \((16\% \ (n = 9) \mathrm{and} \ 5\% \ (n = 3)\) counterparts.

The same subjects who reported trying to lose weight at baseline also reported trying to lose weight at follow-up. Reports of attempted slimming practices included, at mean ages of 11 and 12 years respectively, eating less fatty foods
(78% vs. 66%), doing more exercise (63% vs. 68%) and skipping meals (13% vs. 14%). At baseline and follow-up respectively, 14% ($n = 8$) and 10% ($n = 6$) of underweight and 30% ($n = 26$) and 33% ($n = 26$) of normal-weight subjects reported that they had previously tried to lose weight. Initial and one-year follow-up showed more girls than boys reporting weight-loss attempts (42% ($n = 45$) vs. 20% ($n = 17$) at mean age of 11 years, $P \leq 0.05$; 41% ($n = 41$) vs. 21% ($n = 18$) at mean age of 12 years, $P \leq 0.05$). Weight-loss attempts were comparable among the SEC classifications. Although comparable at baseline, at follow-up overweight girls were more likely to desire a thinner body shape than overweight boys (67% ($n = 20$) vs. 33% ($n = 6$), $P = 0.051$) and twice as likely to report that they had tried to lose weight (70% ($n = 21$) vs. 32% ($n = 6$), $P \leq 0.05$).

**Discussion**

The present study indicates that the prevalence of very overweight young Irish adolescents has increased three-fold since 1990. This trend is indicated by comparing the current sample with a nationally representative sample assessed over 10 years ago using the same criterion for defining overweight. However, the current prevalence of overweight was difficult to quantify according to current standards for assessing overweight in childhood and adolescence. The prevalence varied, ranging from 2 to 11% at baseline and from 5 to 13% at follow-up, within the same group of young adolescents depending upon which cut-off was applied to weight-for-height data. Accordingly, the trend for overweight ranged from zero to 3% over one year, highlighting the difficulties in assessing changes in weight status among young Irish adolescents. Fear of fatness was found to be prevalent among overweight individuals and among girls in particular. Prevalence of unhealthy weight-loss strategies (skipping meals) increased by 1%, to 14%, in this one-year follow-up study.

The limitations involved in the present study in estimating the change in obesity prevalence since 1990 among young Irish adolescents need to be considered. First, the population in the present study was recruited from inner-city Dublin schools whereas the children in the INNS were a nationally representative sample recruited from schools in urban and rural areas throughout the Republic of Ireland. It is possible that children in inner-city Dublin may have experienced greater increases in overweight prevalence compared with those in a nationally representative sample, thereby lessening the observed increase. Comparisons within or between urban and rural populations were not made in the INNS and therefore differences in weight status that may have existed and may still persist remain unknown. However, the recruitment of schoolchildren in the present study was comparable to that employed by the INNS, who stratified schools attended to account for school type, sex mix, socio-economic and age representation. Furthermore, the measurement techniques for weight and height and type of equipment used by trained personnel of the current study were very similar to those employed by trained dietitians in the INNS. Finally, the reasonable response rate to the present study should ensure a good representation of those invited to take part and assuage bias and inflated variance. The criterion used by the INNS to assess overweight, irrespective of age and sex, was defined at as BMI cut-off of 26 kg m$^{-2}$ – this actually corresponds to the cut-off for adults (over the age of 18 years). According to the standard definitions used in this study, it represents the extreme end of overweight status among young adolescents, exceeding the 95th and 98th cut-offs and corresponding with the adult cut-off of 30 kg m$^{-2}$ at age 12 years for both boys and girls. The results from the current study indicate that this group of children has grown over the last decade. However, this
positive trend in obesity prevalence was not apparent after one year of follow-up, according to several weight status assessment methods used in this study, which indicates that longer follow-up periods are necessary to detect secular changes in overweight prevalence among children.

The increasing proportion of overweight children reported in the present study is reflected in the growing prevalence of overweight and obesity documented in other industrialised countries\(^1\).\(^{24}\)\(^{25}\). There are serious health consequences accompanying the rising prevalence of childhood obesity; most notably the increasing trend of paediatric cases of type 2 diabetes\(^26\). However, assessing the trend in weight status among Irish adolescents represents a particular problem. There are different references of BMI-for-age and sex to choose from for the assessment of body weight status in children and adolescents\(^6\)\(^8\)\(^9\). In the present study, cut-offs derived from large samples, and surveys of pooled data collected over a period of time and which employed quality control measures that minimised measurement error, were used to assess BMI data\(^6\)\(^8\)\(^9\). The prevalence of overweight and obesity among this sample is perhaps best represented according to the IOTF cut-offs\(^9\) that provide internationally comparable prevalence rates of overweight and obesity in children that are linked to the widely accepted adult cut-off points of a BMI of 25 and 30 kg m\(^{-2}\). Use of IOTF cut-offs for BMI-for-age thus allows for easier follow-up of paediatric populations through adulthood. But the current IOTF recommendations do not have a cut-off point for underweight since linking to the adult cut-off of 18.5 kg m\(^{-2}\) would yield an unacceptably high proportion of children at risk, while, to date, insufficient childhood morbidity evidence exists to suggest that this is valid\(^9\).

However, since weight-loss attempts have been shown to go hand in hand with fear of fatness, even among underweight girls\(^22\), it is critical to have a criterion of underweight.

A nationally representative sample that reflects secular trends, accounting for economic, nutritional and environmental influences on size and growth potential, is an important consideration\(^8\)\(^,\)\(^27\). Therefore for the assessment of fear of fatness in the present study, ARW based on Irish clinical growth charts\(^10\) was used. However, the limitations of the Irish clinical growth standards have to be considered. Current wisdom regarding the construction of reference growth charts recommends that statistical smoothing procedures should be applied to transform data at each age to normal distribution\(^28\). The Irish growth centiles were based on cross-sectional measurements of weight and height, from approximately 100 boys and 100 girls for each year from age 5 to 19 years inclusive, that were smoothed graphically without sophisticated statistical procedures and used to produce longitudinal standards\(^10\). This methodology falls short of criteria described above for choosing cut-off points of BMI that identify overweight. Furthermore, indices based on relative weight, for example ARW, assume that the amount of body fat is the same irrespective of stature, underestimating the level of obesity in short individuals\(^29\). In this regard, BMI compared with ARW has a significantly stronger correlation with subcutaneous body fat in adolescents, as was found with the triceps skinfold in this study and others\(^39\).

Finally, given that an excessive abdominal adiposity, rather than excessive body mass, is associated with adverse health\(^30\), it is more appropriate to use the indicator that most closely corresponds to risk. Unfortunately, the INNS did not report waist circumference and therefore trends in abdominal fat distribution cannot be assessed. Recent evidence from Great Britain indicates that the increasing trends in waist circumference during the past 10–20 years greatly exceed those in BMI\(^31\). The authors suggest that these trends may represent an accumulation of fat mass that is being obscured by a reduction in muscle mass\(^31\).

Fear of fatness, expressed as a desire to be thinner, was found to relate to weight-loss attempts and to affect more girls, increasing with age, than boys. Furthermore, two-thirds of overweight girls, compared with one-third of overweight boys, reportedly desired a thinner body shape, with twice as many overweight girls than boys trying to lose weight at follow-up. Comparing the desire to be thinner among young adolescent girls in the present study (39%, mean age 12 years) with that reported among older adolescent Dublin girls (59%, mean age 15 years)\(^22\) suggests that negative body image increasingly pervades girls’ progression to womanhood.

Whilst avoidance of obesity is recognised as a healthy practice, inappropriate dieting and weight-loss behaviours known to accompany fear of fatness among adolescents, including purging behaviours and the use of tobacco\(^22\), alcohol and illegal drugs\(^22\), may actually pose a far greater threat to their health than obesity\(^17\). In light of this, public health initiatives that aim to tackle the rising trends in obesity must also tackle the negative body image culture that pervades adolescence, particularly among girls. The most successful interventions are likely to be those that recognise that body image dissatisfaction is influenced by many social and environmental factors, including the media, peers and family\(^32\)\(^,\)\(^35\), and promote environments (for example, in schools and communities) that support active living and healthy eating\(^2\).

A robust and standard method for identification of Irish children and adolescents at risk of overweight is essential for the evaluation of public health strategies that address the serious and growing problem of obesity. The development of health promotion initiatives, such as healthy eating and active living, which prevent overweight among adolescents without exacerbating the fear of fatness, is urgently needed – particularly among older girls and overweight children.
Obesity and fear of fatness among Dublin schoolchildren

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